Title: APOE ε4 is associated to younger age of ischemic stroke onset but not to stroke outcome

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Word count text (excluding figure legend, appendix 1 and references): 1071 (max 750). Character count title (including spaces and punctuation): 87 (max 96). Number of references: 9, Tables: 0, Figures: 1.

Search terms: 1. Ischemic Stroke, 2. Functional outcome, 3. Genetic Association Study, 4. *APOE*, 5. Apolipoprotein E.

Study funding: The Swedish Heart and Lung Foundation (CJ, AL); the Swedish Research Council (CJ); the Swedish Stroke Association (CJ, AL); the Swedish state under the agreement between the Swedish government and the county councils, the ALF agreement (CJ, AL); the Gothenburg Foundation for Neurological Research (CL); the Freemasons Lodge of Instruction EOS in Lund (AL); the Foundation of Färs & Frosta—one of Sparbanken Skåne's ownership Foundations (AL); Lund University (AL); Region Skåne, Skåne University Hospital (AL); the National Health and Medical Research Council, Australia (JM), and the NIH-NINDS (NSR).

Funding for the entire GISCOME study is listed in Söderholm M, Pedersen A, Lorentzen E, et al. Genome-wide association meta-analysis of functional outcome after ischemic stroke. *Neurology* 2019;92:e1271-e1283.

Disclosures

Cecilia Lagging, Erik Lorentzen, Tara M Stanne, Annie Pedersen, Martin Söderholm, John W Cole, Katarina Jood, Robin Lemmens, Chia-Ling Phuah, Natalia S Rost, Vincent Thijs, Daniel Woo and Jane M Maguire report no disclosures relevant to the manuscript. Arne Lindgren reports personal fees for Advisory Board, Speech and Seminar Participation from: Bayer, Astra Zeneca, Boehringer Ingelheim, BMS Pfizer and Reneuron. Christina Jern reports no disclosures relevant to the manuscript.

INTRODUCTION

Stroke outcome is determined by a complex interplay, where age and stroke severity are predominant predictors. Studies on hemorrhagic stroke indicate that *APOE* genotype is a predictor of post-stroke outcomes^{1, 2}, but results from studies on ischemic stroke are more conflicting^{2, 3}. There is one study suggesting an influence of *APOE* genotype on age of ischemic stroke onset⁴, and sex specific effects on outcome have been reported⁵. Taken together, there is a need for larger studies on *APOE* and ischemic stroke outcomes with integrated information on age, severity and sex.

The three common APOE alleles ($\varepsilon 2$, $\varepsilon 3$, $\varepsilon 4$) can be separated by a combination of two single nucleotide polymorphism (SNPs). Thus, associations to APOE alleles are not directly captured in a regular genome-wide association study (GWAS), where each SNP is investigated separately. We derived APOE common alleles and investigated the interplay between APOE, age of ischemic stroke onset, severity, sex and outcome within a large international collaboration; the Genetics of Ischemic Stroke Functional Outcome (GISCOME) network.

METHODS

The design and results of the first GWAS on ischemic stroke outcome within GISCOME have been reported⁶, and the present study comprises the 6,165 cases included in this GWAS. Each center individually obtained appropriate ethical approval and participant consent. Baseline stroke severity was scored by the NIH Stroke Scale (NIHSS) and 3-month functional outcome by the modified Rankin Scale (mRS). Genotyping was performed with SNP arrays with subsequent imputation to the 1000 Genomes Phase 3 reference panel as described⁶. Genotyping of the SNPs rs429358 and rs7412 can distinguish *APOE* alleles £2, £3 and £4. In

the present study, we investigated effects of *APOE* minor alleles $\varepsilon 4$ and $\varepsilon 2$ separately in comparison to the most common allele $\varepsilon 3$. To this end, $\varepsilon 4$ allele count was defined as the continuous imputed minor allele dosage of rs429358(C), excluding samples with minor allele dosage >0.4 for rs7412(T), and vice versa for $\varepsilon 2$, as depicted in Figure 1A. Each cohort was analyzed separately, and for each analysis, cohorts with an effective number of minor alleles ≤ 5 or an extreme effect size ($\beta > 100$) were excluded. Results from the remaining cohorts were meta-analyzed using inverse variance weighted fixed effects, unless there were signs of heterogeneity ($P_{heterogeneity} \leq 0.05$) in which case random effects were used.

We used Directed Acyclic Graphs (DAGs) to investigate associations between *APOE*, age at stroke onset, stroke severity and outcome. A DAG illustrates associations between variables according to a definite direction of causality as depicted by the arrows connecting the variables. For instance, *APOE* can influence age of stroke onset and/or stroke severity, but reverse causality is unlikely as *APOE* genotype is determined at conception. As age and stroke severity are well-established predictors of stroke outcome, we aimed to account for both possible direct effects of *APOE* on outcome and/or indirect effects via associations to age and/or stroke severity as depicted by the three different arrows originating from *APOE* ε4 vs ε3 and ε2 vs ε3 in the DAGs in Figure 1B and D, respectively. Thus, adjustments were made accordingly. In addition, all genetic analyses were adjusted for ancestry (the five first principal components). Prespecified sex-stratified analyses were performed. Associations between allele count, age and stroke severity were analyzed by linear regression. Associations to outcome were analyzed with logistic (dichotomized mRS 0-2 vs 3-6) and ordinal logistic regression.

RESULTS

Increasing allele count of $\varepsilon 4$ was associated with younger age of stroke onset (β - 1.8, P<0.001, Figure 1B). This association was consistent across a majority of cohorts (Figure 1C), significant in both sexes and in cases with first ever stroke only (data not shown). There was an association between $\varepsilon 4$ allele count and favorable outcome (mRS ≤ 2) when adjusting only for ancestry, but this association was no longer retained after additional adjustment for age and stoke severity (Figure 1B).

For $\varepsilon 2$ allele count, we found a direct association to poor outcome (mRS>2) in men after adjustment for ancestry, age and stroke severity (Figure 1D). No such association was detected in the whole sample or in women. Neither $\varepsilon 4$ nor $\varepsilon 2$ allele count showed association to stroke severity.

DISCUSSION

This is the largest meta-analysis with combined information on *APOE* common alleles, age of ischemic stroke onset, severity and outcome to our knowledge. We found that increasing $\varepsilon 4$ allele count was associated with younger age of stroke onset, a finding that is in line with a previous meta-analysis of candidate gene studies⁴. However, we found no evidence of a direct effect of $\varepsilon 4$ on outcome, similar to one recent candidate gene study (n=786)⁷ and one meta-analysis (n=1,453)².

Future studies will have to elucidate the mechanisms behind the association between *APOE* \$\parable{E}\$ allele count and younger age of ischemic stroke onset. As there are well-known associations between *APOE* and lipoproteins, putative mechanisms include effects on lipids and atherosclerosis. In line with this, a large pooled analysis of studies on *APOE* genotype and lipids and other biomarkers, found a clear dose-response association between *APOE* and

low-density lipoprotein cholesterol (LDL) as well as carotid intima media thickness (IMT), with highest LDL and IMT in subjects homozygote for the *APOE* ε4 allele⁸.

In the sex stratified analysis, we found an association between increasing $\varepsilon 2$ allele count and poor outcome in men. To the best of our knowledge, this association is novel. However, sex specific effects of *APOE* on ischemic stroke outcome have previously been reported⁵ and are not unreasonable to assume from a cardiovascular viewpoint. The $\varepsilon 2$ allele has also been associated with increased bleeding volume and poor outcome following lobar intracranial hemorrhage¹. Very speculatively, our results might thus be explained by a higher frequency of hemorrhagic transformation in male $\varepsilon 2$ carriers.

The GISCOME study has the advantage of being the largest sample of genetic and ischemic stroke outcome data available. Study limitations have been previously discussed⁶. In addition, for the sex-stratified analyses in this study our sample size was small. Furthermore, we used imputed values from SNP arrays to establish *APOE* alleles. However, imputed genotyping of these SNPs against the 1000 Genomics reference panel has previously been reported reliable in inferring the three common *APOE* alleles⁹.

In conclusion, this study shows that APOE $\varepsilon 4$ carriers have a younger age of ischemic stroke onset. We also detected worse functional outcome in male $\varepsilon 2$ carriers, a finding that awaits replication. Given these findings, even larger studies would be of interest in order to investigate associations between APOE alleles and ischemic stroke outcomes in different age and sex strata.

FIGURE 1.

Figure title: *APOE* allele distribution and associations to age of ischemic stroke onset, stroke severity and outcome.

Figure legend:

A: Distribution of APOE alleles according to the SNPs rs429358 and rs7412. As the allele counts are inferred from imputation they are given as a continuum between homozygosity for the major allele and homozygosity for the minor allele, with slightly jittered positions in the graph to avoid over-plotting. N=5325 cases were included in the analyses of ε 4 vs ε 3 and N=4519 were included in the analyses of ε 2 vs ε 3. Red positions correspond to cases excluded from both analyses.

B and **D**: Directed Acyclic Graphs (DAGs) displaying associations between *APOE* allele count and age of ischemic stroke onset, stroke severity (baseline NIH Stroke Scale score), and dichotomized 3-month modified Rankin scale score (mRS score 0-2 vs 3-6). N indicates number of cases with non-missing information, except for *APOE* allele count where N refers to maximum number of cases included in the analysis, i.e. cases with allele dosage ≤0.4 for rs7412(T) for ε4 vs ε3, and ≤0.4 for rs429358(C) for ε2 vs ε3. Figure 1B examines associations to ε4 allele count and includes both sexes whereas figure 1D displays associations to ε2 allele count in men only. Associations are reported in the squared text boxes as β, *P* value derived from linear regression for associations to age and stroke severity, and odds ratios (OR), *P* value derived from logistic regression for associations to poor outcome (mRS>2). Adjustments are indicated in the parentheses as follows: PC, adjusted for ancestry (the five first principal components); A, age adjusted; S, stroke severity (baseline NIHSS) adjusted. *Refers to result from random effects meta-analysis. Arrow thickness illustrates standardized effect size after the full adjustment specified in the respective text box.

Arrow color refers to the direction of the effect. A dotted arrow indicates a non-significant association.

C: Bubble chart showing median age of ischemic stroke onset and $\epsilon 4$ allele frequency for individual cohorts in GISCOME. The cohorts are described in Söderholm M, Pedersen A, Lorentzen E, et al⁶. Bubble diameter is proportional to the number of cases. Bubble color refers to the effect size (β) of $\epsilon 4$ on age of stroke onset derived from linear regression.

APPENDIX 1 – list of authors

Role – "Author" for all listed below.

Name	Location	Contribution
Cecilia Lagging,	University of Gothenburg,	Interpretation of the data, drafting the
MD	Gothenburg, Sweden	manuscript for intellectual content
	Sahlgrenska University	
	Hospital, Gothenburg, Sweden	
Erik Lorentzen,	University of Gothenburg,	Analysis and interpretation of the data,
MSc	Gothenburg, Sweden	drafting parts of the manuscript for intellectual
		content
Tara M Stanne,	University of Gothenburg,	Interpretation of the data, revising the
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Annie Pedersen,	University of Gothenburg,	Interpretation of the data, revising the
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Martin Söderholm,	Lund University, Lund, Sweden	Interpretation of the data, revising the
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Robin Lemmens,	University of Leuven; VIB,	Major role in the acquisition of data, revising
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Jane M Maguire,	University of Technology	Design and conceptualization of the
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Arne Lindgren,	Lund University, Lund, Sweden	Design and conceptualization of the
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	University of Gothenburg,	Design and conceptualization of the
Christina Jern,	Gothenburg, Sweden	GISCOME and the present study, major role
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$\label{eq:APPENDIX 2-list of co-investigators} APPENDIX\ 2-list\ of\ co-investigators$

Role – "Co-investigator in the Genetics of Ischemic Stroke Functional Outcome (GISCOME) network" for all listed below.

Name	Location	Contribution
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Worrall, MD	Charlottesville	GISCOME, major role in the acquisition of
		data

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