

Asset impairments surrounding CFO turnover

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Abstract

We examine the effect of Chief Financial Officer (CFO) turnover on the likelihood and size of asset impairments. Using Australian data, we find evidence consistent with incoming CFOs recording larger asset impairments, particularly when they are external hires with prior listed experience, or receive equity-based compensation. Our results also indicate that outgoing CFOs report higher profits in the year prior to their departure by reporting fewer and smaller asset impairments. This effect is larger when outgoing CFOs move to another listed firm or receive equity-based compensation, consistent with reputational and compensation incentives to maximise earnings. Overall, our results provide evidence of the independent influence CFOs have on financial reporting.

JEL: M41, G34, M40

KEYWORDS: CFO turnover; asset impairment; earnings management+

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1. Introduction

The objective of this study is to extend prior research on asset impairments (Bond et al., 2016) by examining the effect of Chief Financial Officer (CFO) turnover on the *likelihood* and *size* of non-current asset impairments, including both goodwill and other assets, in the period prior to, during, and immediately after CFO turnover.¹ Despite evidence regarding the significance of CFO turnover on firms' reported results (e.g., Bamber et al., 2010; Brochet et al., 2011; Dejong and Ling, 2013; Demerjian et al., 2013; Ge et al., 2011; Geiger and North, 2006), whether outgoing and incoming CFOs use asset write-offs as a vehicle to manipulate accounting profit remains unexplored.² This lack of research is surprising given that the evaluation of the recoverability of non-current asset carrying values provides substantial discretion for CFOs to influence firms' accounting numbers.

The motivation for this study is threefold. First, prior research largely attributes the “big bath” phenomenon to a change in a firm's Chief Executive Officer (CEO) (e.g., Moore, 1973; Murphy and Zimmerman, 1993; Pourciau 1993; Strong and Meyer, 1987; Wells, 2002). CEO turnover is, therefore, considered to be an important determinant of asset impairments (e.g., Bond et al., 2016). However, the primary responsibility for decisions involving sophisticated financial judgement is delegated to the CFO, and it is widely acknowledged that CFOs influence a firm's reported results.³ Prior evidence also suggests that CFO career opportunities are hindered by the arrival of a new CEO, increasing the likelihood of CFO replacement following CEO turnover (Collins et al., 2009; Fee and Hadlock, 2004; Mian, 2001). Therefore, the “big bath” phenomena may be partly driven by concurrent CEO-CFO appointments and/or CFO turnover alone, which have been overlooked in prior research.

The second motivation for this study is to provide evidence on whether CFOs serve as “monitors” of financial reporting integrity or act in self-interest around their turnover. This examination is important for firm stakeholders, as CFOs are legally expected to uphold financial reporting integrity and suffer labour market penalties for financial reporting failures

¹ In this study the terms “non-current asset impairments”; “asset write-offs”; and “asset write-downs” are used interchangeably to refer to the impairment of non-financial assets such as goodwill, identifiable intangibles and property, plant and equipment under AASB 136 *Impairment of Assets*.

² As far as can be ascertained the only study examining reductions in reported earnings by incoming CFOs is Geiger and North (2006) who focus only on discretionary accruals.

³ For example, Aier et al., (2005); Barua et al., (2010); Bedard et al., (2014); Collins et al., (2009); Engel et al., (2015); Feng et al., (2011); Ge et al., (2011); Geiger and North, (2006); Jiang et al., (2010); Kim et al., (2011); Wang et al., (2012); and Wang, (2010).

(Bedard et al., 2014; Collins et al., 2009; Engel et al., 2015; Haislip et al., 2015). However, *outgoing* CFOs may act in self-interest due to compensation and career concerns, aiming to increase earnings in the year prior to their replacement. Specifically, departing CFOs have incentives to achieve performance targets in order to maximise their final pay (e.g., Geiger and North, 2006), maintain their position within the firm (e.g., Fee and Hadlock, 2004; Feng et al., 2011; Mian, 2001; Park et al., 2014) and increase the value of their human capital in the managerial labour market (e.g., Brickley et al., 1999; Fama, 1980; Mian, 2001). Similarly, *incoming* CFOs may engage in “big bath” accounting in their year of appointment (e.g., Geiger and North, 2006) attributing poor results to their predecessors.

The final motivation for this study is to extend limited prior research on the factors associated with the recognition of non-current asset impairments. Whilst Bond et al., (2016) find that impairment decisions appear to be discretionary, and Vanza et al., (2018) document asset impairment recognition is not associated with information asymmetry or the uncertainty about future returns, there is currently no evidence on whether non-accounting factors such as CFOs and executive turnover are associated with impairment decisions.

Our analysis is conducted using an Australian sample of 463 CFO changes over the period 1 July 2007 to 30 June 2015. We focus on firm-year observations around CFO turnover to limit the need to control for the endogenous nature of CFO turnover events.⁴ We also investigate all asset impairments as opposed to just goodwill impairments, as Bond et al., (2016) find that non-goodwill asset impairments are dominant in the Australian environment. We find that *incoming* CFOs initiate significantly larger asset impairments in the year of appointment.⁵ Conversely, we find that prior to their departure, *outgoing* CFOs are significantly less likely to record an asset impairment, and if they record one, it is significantly smaller in size.

Prior research indicates that the incentives and ability to manipulate earnings varies with the circumstances of the appointment (Geiger and North, 2006; Murphy and Zimmerman 1993; Pourciau 1993; Wells, 2002), therefore we partition our sample based on various characteristics associated with the CFO turnover event. We document that when the outgoing CFO continues employment with the firm in another capacity, there is no evidence of impairments around the turnover. In contrast, when the outgoing CFO does not remain with the firm, obtains

⁴ In additional testing (Table 9) we test if asset impairments are greater using an expanded sample that includes firm-years without CFO turnover.

⁵ We control for CEO turnover in all empirical tests.

employment with another listed firm, or has equity incentives there is a decrease in the likelihood and size of asset impairments prior to their departure. In terms of incoming CFOs, we document an increase in the size of impairments in their appointment year when they are an external hire, have prior experience with a listed firm, or receive equity compensation. Internally recruited CFOs, however, appear to wait until the year after their appointment to record asset impairments.

Additional testing demonstrates that CFOs exercise significant independent influence on firms' financial reporting. Specifically, when we compare subsamples of firm-years with CEO turnover to those with CFO turnover, the likelihood and size of non-current asset impairments does not differ across the two groups. We also find, after dropping years with CEO turnover, that the likelihood and size of asset impairment increases in the year following a CFO turnover. The evidence in this study is therefore similar to Geiger and North (2006), who document incoming CFOs recognize income decreasing discretionary accruals in the year following a CFO change. The results also indicate that CFOs exert independent influence on financial reporting decisions.

This study makes a number of important contributions. First, it extends the existing literature on "big bath" accounting by emphasising the significance of CFO turnover on firm financial reporting (e.g., Dejong and Ling, 2013; Demerjian et al., 2013; Ge et al., 2011; Geiger and North, 2006). While the majority of the literature focuses on CEO turnover and documents "big bath" accounting in the year of a new CEO's appointment, our study highlights the independent role of CFOs in earnings management. We demonstrate that the judgement required in estimating the recoverability of non-current asset carrying values provides opportunities for earnings manipulation by CFOs, irrespective of concurrent CEO turnover. Specifically, we show that CFOs can independently manage earnings through their decisions to recognise asset write-offs, with the size and timing of these write-offs appearing to align with their self-interest. This finding underscores the influence of CFO turnover on financial reporting practices and informs stakeholders about the potential for opportunistic behaviour by CFOs during executive transitions.

Second, whilst goodwill impairments have been studied extensively in prior research (Hayn and Hughes 2006; Li et al., 2011), there is currently limited literature on the determinants of non-current asset impairments (Bond et al., 2016; Vanza, et al., 2018). We contribute to and

extend this literature by examining whether executive turnover is an additional factor influencing asset impairment decisions. Our results indicate that both CFO and CEO turnover independently drive asset impairments, suggesting that these factors should be included as controls in studies examining the recognition of non-current asset impairments. This inclusion enhances the understanding of the determinants of asset impairments and provides a more comprehensive view of the factors influencing financial reporting.

Third, our findings have implications for regulators (e.g., ASIC), auditors, and audit committees by suggesting that heightened scrutiny is needed for firms making CFO changes, as our evidence indicates these impairments are not justified by firm fundamentals. Specifically, ASIC may consider increasing surveillance of firms experiencing CFO turnover to identify unjustified asset impairments. Similarly, additional audit effort may be necessary around CFO turnover events to assess the validity of asset impairment decisions. Increasing the scrutiny of asset impairments is important for shareholders, as prior research highlights the negative market reaction to asset write-offs (Riedl 2004; Bens et al., 2011).

The remainder of this study is structured as follows. Section 2 reviews prior literature and develops hypotheses. The regression models, which estimate the effect of CFO turnover on the likelihood and size of non-current asset impairments, are detailed in Section 3. Section 4 describes the sample selection process, whilst the empirical results are discussed in Section 5 along with additional analyses. Finally, Section 6 provides concluding remarks.

2. Literature Review and Hypotheses Development

It is widely acknowledged that the CFO is the leader of the finance and accounting function. The growing importance of CFOs' stewardship responsibilities is reflected in the US Sarbanes-Oxley Act and similar Australian regulatory reforms (i.e., CLERP 9), which require both the CEO and CFO to personally certify the accuracy and completeness of financial information. Following this legislative elevation of CFO financial oversight responsibility, an emerging literature recognises the rise of CFO visibility, power, and importance beyond that of other executives (e.g., Aier et al., 2005; Bedard et al., 2014; Wang, 2010; Wang et al., 2012). As top-level managers with the capacity to exert control over the financial reporting process, CFO turnover is expected to be an important determinant of financial reporting integrity (e.g., Bamber et al., 2010; Brochet, et al., 2011; Dejong and Ling, 2013; Demerjian et al., 2013; Ge et al., 2011; Geiger and North, 2006). The ability of CFOs to independently influence a firm's

accounting choices is particularly pertinent in complex financial reporting areas such as asset impairment, which requires the exercise of substantial judgement, professional skepticism, and discretion.

2.1 *Outgoing CFOs' Financial Reporting Incentives*

CFOs may be motivated to maximize reported income prior to their departure for several reasons. First, compensation incentives may motivate CFOs to engage in earnings management particularly in the case of routine retirement (i.e., the horizon problem). It is well-established that bonus-based compensation contingent on financial measures provides top level management with motivations to manipulate income (e.g., Healy, 1985) and prior research has found that CFO equity awards are positively associated with accruals management (Jiang et al., 2010), the likelihood of beating analyst forecasts (Jiang et al., 2010), the manipulation of information flow to capital market participants (Kim et al., 2011), and insider trading (Wang et al., 2012). Therefore, CFOs may use accounting discretion to reach performance targets to maximize their pay in the year preceding their departure (e.g., Geiger and North, 2006). Second, CFOs are also susceptible to career and reputational pressure to avoid reporting poor performance (e.g., Engel et al., 2015; Graham et al., 2005; Mian, 2001). For example, when CFO turnover is disciplinary, it is preceded by weak stock price (Mian, 2001) and/or operating performance (Engel et al., 2015; Mian, 2001). CFOs dismissed in these circumstances are also penalised in the labour market (Fee and Hadlock, 2004).⁶ Therefore, outgoing CFOs have incentives to use their accounting discretion to maximise firm performance.

These career, reputation and compensation payoffs potentially serve as significant motivating factors in the financial reporting decisions of *outgoing* CFOs. In summary, *outgoing* CFOs are expected to bias earnings upward in the year prior to departure leading to a lower likelihood and size of non-current asset impairments.

2.2 *Incoming CFOs' Financial Reporting Incentives*

It has long been recognised that a change in leadership provides an incentive for *incoming* CEOs to bias earnings downward (e.g., through large asset write-offs) in the initial year of appointment. This poor firm performance is then blamed on previous management and the subsequent improvements in performance (paper profits from accrual reversals) are attributed

⁶ Fee and Hadlock (2004) find that displaced non-CEO executives who do secure new employment typically take positions that are inferior (i.e., based on firm size, compensation and seniority) to their prior employment.

to the new management. This behaviour is what the literature describes as an “earnings bath” (e.g., Moore, 1973; Murphy and Zimmerman, 1993; Pourciau 1993; Strong and Meyer, 1987; Wells, 2002). It is plausible that incoming CFOs engage in similar behaviours and Geiger and North (2006) use discretionary accruals to provide evidence consistent with this rationale.⁷

The discussion above suggests that the likelihood and size of impairments are expected to decrease/(increase) in the period before/(after) CFO turnover. This leads to Hypothesis One:

H1(a): There is a decrease in the likelihood and size of non-current asset impairments in the period prior to CFO turnover.

H1(b): There is an increase in the size and likelihood of non-current asset impairments in the period(s) at or after CFO turnover.

2.3 The Nature of CFO Turnover and Impairment Decisions

Prior research investigating the effect of new executives on financial reporting identifies that the impetus and opportunities for earnings manipulation vary with the circumstances of the appointment. For example, the results in Geiger and North (2006) are predominantly driven by externally appointed CFOs, consistent with outsiders bringing in new perspectives and having less commitment to the status quo (e.g., Büttner et al., 2013; Geiger and North, 2006; Karaevli, 2007). Conversely, internal successions convey continuity in leadership dependent on existing practices, firm-specific knowledge, and established networks within the firm (Büttner et al., 2013; Karaevli, 2007). Given these findings, we also investigate if certain aspects of CFO turnover are driving impairment decisions. This is described in section 4 below.

3. Research Design

The model below, based on Bond et al., (2016), is estimated to test the association between CFO turnover (*CFOTURN*) and the likelihood (*IMPAIR*) and size (*IMPAIRSIZE*) of non-current asset impairments:⁸

⁷ It should be noted that unlike CEO turnover research which provides consistent support for the “big bath” in the initial reporting period of the CEO appointment (i.e., t), Geiger and North (2006) document discretionary reductions in earnings in the year subsequent to a CFO appointment (i.e., $t+1$).

⁸ Bond et al., (2016) report that only 27.4 percent of firms impair goodwill, and most asset impairments relate to tangible assets. Furthermore, of the firms recognizing impairment but not goodwill impairment, 61.4 percent of firms had goodwill on the balance sheet. This supports the focus in this study on impairments of non-current assets more generally.

$$\begin{aligned}
\text{IMPAIR/IMPAIRSIZE} = & \beta_0 + \beta_1\text{CFOTURN} + \beta_2\text{CEOTURNOVER} + \beta_3\text{BM} + \beta_4\text{YRS} + \beta_5\text{BHR} \\
& + \beta_6\text{EARN} + \beta_7\text{CF} + \beta_8\text{GOV_NED} + \beta_9\text{BIG4} + \beta_{10}\text{LNMVE} + i.\text{ind} + \varepsilon
\end{aligned}
\tag{1}$$

Model (1) is estimated (*CFOTURN*) separately for the year prior (*CFOTURN_PY*), the year of (*CFOTURN*) and following CFO turnover (*CFOTURN_FY*), consistent with Bond et al., (2016). *IMPAIR* is defined as a binary variable coded as one if a non-current asset impairment is recognised, zero otherwise. *IMPAIRSIZE* represents the total non-current asset impairment loss (excluding any impairment reversals) recognised by the firm in the current financial year scaled by prior year total assets.⁹ Following Bond et al., (2016), model (1) includes controls for the factors that AASB 136 identifies as potential indicators of impairment. A book value in excess of market value is an external indicator of impairment (AASB 136, paragraph 12(d)) and therefore is included as an independent variable (*BM*) and measured as the ratio of the book value of equity (adjusted for the recognition of asset impairments) to the market value of equity at the end of the prior financial year. As a decline in market value is potentially not temporary for firms where $BM > 1$ or $BM < 0$ for more than one year, the model also includes a binary variable coded as one if $BM > 1$ or $BM < 0$ for two consecutive reporting periods prior to the current financial year end, zero otherwise (*YRS*). Both *BM* and *YRS* are expected to have a positive relation to *IMPAIR* and *IMPAIRSIZE*. Consistent with significant declines in market value being an external indicator of impairment (AASB 136, paragraph 12(a)), model (1) also controls for the buy-hold return of the stock over the prior financial year (*BHR*). A negative relation between *BHR* and the likelihood and size of impairments is expected. Given prior evidence we also add a control for CEO turnover (Wells, 2002) (*CEOTURNOVER*).

Model (1) also includes net profit after tax per share adjusted for impairment charges (*EARN*), and aggregate cash flow from operating and investing activities per share (*CF*). These variables control for internal indicators of impairment (AASB 136, paragraph 14(a)-(d)) and are expected to be negatively related to *IMPAIR* and *IMPAIRSIZE*. To control for effective corporate governance and external monitoring, the model includes controls for board independence (*GOV_NED*) and the use of a Big 4 audit firm (*BIG4*). Finally, the natural

⁹ Non-current asset impairments relate to property, plant and equipment, goodwill and other intangibles, including the impairment of mining assets under construction/development which are being depreciated. The impairment of investments, financial assets, exploration and evaluation assets and mining tenements are excluded. These losses capture any uncertainty regarding the recoverability of economic benefits embodied in the specific asset. Hence, they are not in response to internal and external indicators of impairment based on the performance of the entity as a whole.

logarithm of market capitalization (*LN**MVE*) is included to control for firm size and industry fixed-effects are included to control for factors that vary across industry. Appendix A provides definitions of all variables.

4. Sample selection and characteristics of CFO turnover

The Connect 4 Boardroom database is used to identify 2,602 CFO appointments by ASX listed entities between 1 July 2007 and 30 June 2015.¹⁰ We exclude 701 observations relating to delisted companies, entities in administration and companies newly listed in the year of or prior to a CFO appointment. To maintain uniformity surrounding the timing of disclosures and impairments assessments, 337 CFO appointments in firms with financial reporting periods ending on a date other than 30 June are removed. Lu et al., (2013) document that over 80% of Australian firms have a June year end, and there are systematic differences between firms with June and non-June year-ends. The exclusion of non-June year end firms thus reduces the likelihood that such differences impact our analysis. Also, the uniformity in financial year-end ensures all sample firms are assessing impairment at the same time and based on identical market conditions.

A further 243 observations are excluded due to industry or firm specific financial reporting issues. That is, entities that revalue items of property, plant and equipment, as well as firms in the agriculture and real-estate investment sectors that may not recognize declines in asset values as impairments due to the application of fair value accounting. Financial and utilities firms operate in a different regulatory environment with distinct reporting requirements and are therefore also excluded.

In 241 cases, the CFO appointment is merely a position change for the incumbent CFO who absorbs the Company Secretary function, hence these observations are removed. A further 449 firms are deleted where multiple CFO appointments occur within the same financial year (year t) or recurrent CFO appointments in consecutive financial years. In 139 cases the incoming CFO ceased employment within the appointment financial year (t) or the following financial year ($t+1$) and are therefore excluded. Finally, 29 firms with CFO changes that are only one financial year apart are also removed. These exclusions result in a final sample of 463 unique CFO changes. Data is collected for three financial years surrounding each CFO turnover

¹⁰ This period follows the transition to IFRS by Australian companies in 2005. The sample period commences in the 2008 financial year to ensure comparability and consistent application of accounting policies in the financial year prior to the CFO appointment (i.e., $t-1$) and in variables, which capture the year-to-year change in financial data.

event (i.e., the period prior to $(t-1)$, during (t) and immediately following $(t+1)$ the CFO appointment) resulting in a sample of 1,389 firm-years. A summary of the sample selection process is documented in Table 1.

[Insert Table 1 about here]

Table 2 documents the reason for CFO departures in cases where available information sources (e.g., annual reports, ASX announcements and news articles) provide relevant commentary. There are 17 CFO changes that follow a formal restructuring of the board or senior executive leadership team. In only 5 cases (i.e., 15 firm-years) was the reason for CFO departure consistently described as a retirement and the outgoing CFO was above (or close to) the natural retirement age of 65. Other departures were due to ill health or death (3 events), family related (6 events) or because the former CFO was overlooked for the CEO position (2 events). The reluctance on the part of the directors to candidly discuss the reason for CFO changes is recognised in prior literature (e.g., Park et al., 2014). Therefore, it is unsurprising that in most cases (430 events), disclosures regarding the reason for CFO turnover are absent.

[Insert Table 2 about here]

To address this disclosure deficiency and identify characteristics of CFO turnover, we conduct an information search of ASX announcements, news articles, financial statement disclosures and the executives' LinkedIn profiles. Research into Australian CFOs of ASX 100 listed firms shows that the ratio of internally and externally hired CFOs has been near identical over several years: 52% are hired externally and 48% are hired from within the firm (Chiswick, 2018; 2022). Therefore, an external hire alone is likely not enough to imply a suspicious turnover event which may lead to subsequent write-downs (Vafeas and Vlittis, 2015; Büttner et al., 2013; Mian, 2001). As such, we hand collect the following information which may affect the propensity to record an impairment surrounding CFO turnover: (i) qualifications and prior experience of the incoming CFO, (ii) if the incoming CFO is awarded equity-based compensation, and (iii) if the outgoing CFO continues with the firm in a different capacity or moves to another listed firm.

First, firms typically bring in 'bigger and better' talent, such as a CFO from a listed company, when they wish to overhaul a negative image or improve processes significantly, which could result in impairments. This view is supported by facts showing that 53% of CFO appointments in Australian ASX 100 firms are first-time group CFOs and only 32% of

appointed CFOs have prior public company CFO experience (Chiswick 2022). Second, if the incoming CFO is awarded equity-based compensation it provides an incentive to make write-downs that will result in greater subsequent performance. Third, career progression and promotion are positive signals of CFO quality consistent with ex-post settling up in the labour market (Haislip et al., 2015). Therefore, if the outgoing CFO is appointed to another listed firm, there is likely an incentive to record fewer write-downs prior to departure.

Panel C of Table 3 provides a breakdown of the above characteristics of CFO turnover.

[Insert Table 3 about here]

Despite our advanced search, we were unable to identify all characteristics for every CFO turnover event due to lack of disclosure, therefore the number of observations differs across characteristics. Within our sample, most CFO appointments are from candidates external to the firm (78%) (*I_CFO_EXTERNAL*) and more than half have prior experience with listed firms (66%) (*I_CFO_PRIOR_LISTED*) which is higher than recent research from the ASX 100 (Chiswick, 2018; 2022). Not surprisingly, most are qualified accountants (87%) (*I_CFO_QUALIFIED*) and receive equity-based compensation as part of their remuneration package (75%) (*I_CFO_EQUITY*). In terms of outgoing CFOs, roughly 22% continue with the firm in some capacity (*O_CFO_CONTINUE*) (e.g., as a non-executive director) and 69% take up a position on another publicly listed firm (*O_CFO_NEW_LISTED*). The majority are also qualified accountants (82%) and receive equity-based compensation (79%).

5. Empirical Results

5.1 Descriptive Statistics

Director and other governance data required to estimate model (1) are obtained from the SIRCA Corporate Governance database. Financial data other than asset impairments are obtained from the Morningstar DatAnalysis Premium database. Asset impairments are hand collected from annual reports. Appendix A provides detailed variable definitions. All continuous variables are winsorised at the 5th and 95th percentile to reduce the influence of outliers.¹¹ Table 3 presents descriptive statistics for the variables used in the analysis for the full sample. Mian (2001) documents that CFO turnover is preceded by abnormally high CEO turnover. It is, therefore, unsurprising that CEO turnover is observed in roughly 16 percent of

¹¹ Results are consistent with those presented when the raw data (not winsorised) is used in the analysis.

the sample (*CEOTURNOVER*). Asset impairments are recognised in 27 percent of firm-years (*IMPAIR*) and represent 1.5 percent of total prior period assets (*IMPAIRSIZE*) on average.

[Insert Table 3 about here]

Significant variation across the financial variables is observed. The average market capitalization is \$470.7 million with the median being \$39.1 million. The mean (median) *BM* is 1.041 (0.697) which indicates skewness in the distribution of this variable. An indicator of impairment is present in a significant percentage (36.2 percent) of firm-years, which have a $BM > 1$ or $BM < 0$. Furthermore, 19.3 percent of firm-years have had a $BM > 1$ or $BM < 0$ for two consecutive prior financial reporting periods (*YRS*). Substantial skewness in both market (*BHR*) and accounting (*EARN*) returns is also evident with negative values apparent at the 50th percentile. The mean (median) value of *BHR* is 0.073 (-0.106). *EARN* has a mean (median) value of 0.071 (-0.004). A negative value of *CF* is also evident at the median (-0.009), the mean being 0.022. A material percentage of firm-years, therefore, exhibit both internal and external indicators of impairment. Big 4 accounting firms audit 43.7 percent of the sample (*BIG4*) and 84.2 percent of firm-years have a majority of non-executive directors on the board (*GOV_NED*).

Table 4 presents a comparison of means for the years $t-1$, t and $t+1$ surrounding CFO turnover. A greater percentage of firms recognize asset impairments in year t (30 percent) and $t+1$ (29.4 percent) compared to $t-1$ (21.6 percent). Larger asset impairments are also recognised in year t (1.9 percent of total prior year assets) and $t+1$ (1.6 percent of total prior year assets) versus $t-1$ (1.1 percent of total prior year assets). The differences are statistically significant at the 1% level.

[Insert Table 4 about here]

Table 4 indicates that indicators of impairment may be present in year t and $t+1$ which exhibit a statistically significant lower/(higher) *BHR*/(*BM*) than $t-1$. There is also evidence that a $BM > 1$ or $BM < 0$ has persisted for two consecutive prior financial reporting periods (*YRS*) for a larger percentage of firms in year $t+1$ (21.8 percent) compared to $t-1$ (17.2 percent). A higher *GOV_NED* in year $t+1$ (86.0 percent) versus $t-1$ (81.9 percent) suggests that CFO appointments are accompanied by changes in board composition. Overall, the absence of statistically significant differences across most of the control variables indicates that, year to year, firm characteristics are relatively similar.

5.2 Likelihood and Size of Non-Current Asset Impairments around CFO Turnover

Table 5 presents the results of estimating regression model (1) for the full sample of CFO turnover events.

[Insert Table 5 about here]

Columns (1) to (3) present results for the *likelihood* of non-current asset impairments and columns (4) to (6) present results for the size of asset impairment. Given that *IMPAIR* is a binary variable, a random-effects panel logistic regression is used and odds ratios are reported along with robust z-statistics in parentheses. The sign of the z-statistic indicates whether the predictor variable is associated with higher or lower odds of the outcome occurring.¹² For *IMPAIRSIZE* a Tobit random-effects model is used.¹³

The results in column (1) indicate that firms are 28.3 percent less likely to recognise asset impairments in the year prior to a CFO change (*CFOTURN_PY*) (odds ratio = 0.717, $p < 0.05$). Meanwhile column (4) reports that the value of asset impairments recognised in this period are also lower ($\beta = -0.062$, $p < 0.05$). While there is no evidence that the recognition of asset impairments is more probable in the year of or following a new CFO appointment compared to other years surrounding CFO turnover, column (5) reports that the size of asset write-downs increases by 6.3 percent of total assets in the year of a new CFO appointment (*CFOTURN*) ($\beta = 0.063$, $p < 0.05$). Overall, the findings are consistent with Hypothesis 1(a) and provide partial support for hypothesis H1(b).

Turning to the control variables, CEO turnover is also associated with a higher likelihood and size of asset impairment consistent with prior literature (e.g., Moore, 1973; Murphy and Zimmerman, 1993; Pourciau, 1993; Strong and Meyer, 1987; Wells, 2002). As expected, a significant positive association is found between *BM* and the likelihood (*IMPAIR*) and size (*IMPAIRSIZE*) of asset impairments, while *BHR* significantly reduces the likelihood of asset impairments but not the size. Larger firms (*LNMVE*) and those audited by one of the Big 4 accounting firms (*BIG4*) are also more likely to recognize (and record higher) asset

¹² Due to the number of observations (687 firm-years) that have no variation in *IMPAIR*, it is problematic to estimate a fixed-effects model.

¹³ In Stata *xttobit* fits random-effects Tobit models. Tobit is a non-linear function and the likelihood estimator for fixed-effects is biased and inconsistent.

impairments. In addition, a significant negative association is observed between firm profitability and asset impairment size.

In assessing the results in Table 5, it is important to consider whether CFOs' accounting choices are justifiable by indicators of financial impairment. In separate analysis (not tabulated), the CFO turnover variable is interacted with each of the variables capturing indicators of impairment. Consistent with "big bath" theory, there is no evidence that the likelihood or size of asset impairments in the years surrounding CFO turnover are in response to indicators of impairment. Instead, the only significant results are that in year $t+1$ the likelihood and size of asset impairments decreases in the presence of a $BM > 1$ or $BM < 0$ that has persisted for two consecutive years, but increases with higher earnings.¹⁴

5.3 Characteristics of CFO turnover

We next examine whether characteristics of CFO turnover impact on the impairment decision. To do so, we estimate model (1) on subsamples of firms depending upon the characteristic of the outgoing or incoming CFO. The results are summarised in Table 6.

[Insert Table 6 about here]

We include all control variables and fixed effects as previously discussed and displayed in Table 5 but report only the coefficients, odds ratios and z-stats of the main test variables for brevity. Panel A displays the results for subsamples based on characteristics surrounding the outgoing CFO. In cases where the CFO does not continue with the firm in some capacity, moves on to another listed firm or received equity pay, the likelihood and size of non-current asset impairments are all negative and significantly associated with the year prior to CFO turnover ($CFOTURN_PY$). Intuitively, this makes sense since departing CFOs have reputational and compensation incentives to maximise performance prior to departure, consistent with Hypothesis 1(a). These results are also consistent with those reported in Table 5. All other years remain insignificant in relation to outgoing CFOs apart from when the CFO goes to another listed firm. In this instance, the likelihood (odds ratio 1.899) and size ($\beta=0.039$)

¹⁴ The insignificance of the interaction terms is largely due to issues in predicting asset impairments as noted in Bond et al. (2016).

of non-current asset impairments are higher in the year following a new CFO appointment (*CFOTURN_FY*).¹⁵

Panel B displays results for characteristics surrounding the incoming CFO. In contrast to panel A, there are no significant associations for the year prior to turnover (*CFOTURN_PY*) except when the incoming CFO is an external hire, in which case the size of impairment is smaller ($\beta=-0.044$, $p<0.10$). These results suggest that the impairment decision in the year prior to CFO turnover is influenced by the incumbent CFO at the time and not the incoming CFO. In contrast, when the CFO is an internal replacement, there is a greater likelihood (odds ratio 1.869) and size ($\beta=0.020$) of impairment in the year following but not the year of CFO turnover (*CFOTURN_FY*). This finding is consistent with internal successors requiring time to establish authority over financial reporting. Another possible explanation is internal CFOs not wanting to blame write-offs on a predecessor who was likely a mentor. When the incoming CFO is an external replacement, has prior listed experience or receives equity-based compensation, the size of non-current asset impairments is higher in the year of appointment. These results are consistent with hypothesis 1(b) and the results in Table 5.

5.4 Additional testing and robustness

5.4.1 Classifying turnovers into routine versus non-routine

As additional analyses, we consider all factors surrounding CFO turnover to distinguish routine versus non-routine CFO changes. Prior literature has generally classified CFO turnover as routine if the outgoing CFO has continuing involvement with the firm (i.e., in another executive position or as a member of the board of directors) or an internal replacement is hired (Geiger and North, 2006; Wells, 2002). Therefore, we begin by classifying CFO turnover as ‘routine’ if the above two criteria are met. Second, we consider if the outgoing CFO is appointed to another listed firm since career progression and promotion is a positive signal of CFO quality (Haislip et al. 2015). Third, any CFO turnovers that do not meet the above criteria are classified as routine if there is an internal replacement and non-routine if there is an external replacement or the former CFO cannot be identified. Using this criteria, 206 (44%) of the 463

¹⁵ When testing subsamples of qualified incoming (*I_CFO_QUALIFIED*) or outgoing CFOs (*O_CFO_QUALIFIED*) we find that the likelihood and size of non-current asset impairments are lower in the year prior to turnover regardless of whether the CFO was qualified or not. This result is consistent with our main finding displayed in table 5 and is not surprising given such little variation exists within these variables.

CFO turnover events are classified as routine. Table 7 reports the results of estimating model (1) on subsamples based on this classification.

[Insert Table 7 about here]

Panel A displays results for the routine subsample and Panel B displays results for the non-routine subsample. The results for routine CFO appointments are largely consistent with those documented in Table 6 regarding internal CFO appointments and if the outgoing CFO goes to another listed firm, both of which are reflected in the routine classification criteria. Specifically, the size of impairment is smaller in the year prior to turnover (column 4) and larger in the year following turnover (column 6). BM is positive and significant across all columns.

In regard to non-routine CFO turnover, Panel B shows that the size of asset impairments is larger in the year of appointment, reflecting the results documented in Panel B of Table 6. The controls for CEO turnover and BM are positive and significant across all columns. Overall, our results are consistent with our main findings using this holistic approach to classifying CFO turnover into routine versus non-routine. However, unlike CEO turnover, the paucity in disclosure regarding CFO turnovers means that this method of classifying turnovers into routine versus nonroutine is particularly noisy. Therefore, we believe that investigating the characteristics of outgoing and incoming CFOs separately (as displayed in Table 6) provides greater insight into the reasons behind impairment decisions than subjectively classifying CFO turnovers into routine and non-routine.

5.4.2 Effect of concurrent appointments

Prior research documents that CEO turnover is often accompanied by a “big-bath” in the year of appointment (Wells, 2002). As a result, in our main findings we control for CEO turnover that occurs prior to, during or following the CFO turnover. To provide additional insight, we compare the size of impairments surrounding CFO, concurrent CFO/CEO and concurrent CFO/CEO/Chair turnover to years in which only the CEO changes. The results are displayed in Table 8.

[Insert Table 8 about here]

The average size of impairment is not statistically different between subsamples of CEO turnover and CFO turnover, implying that CFOs have just as much influence over the size of impairment as CEOs. When examining concurrent CFO and CEO turnover, the mean size of impairment is smaller in year $t-1$, which could be driven by motives to maximise performance

prior to departure. Conversely, the mean size of impairment is larger in the year prior to concurrent CFO, CEO and Chair turnover (column 4). Firms generally restructure the senior leadership team following poor firm performance (Mian, 2001) or events which signal financial reporting failures (Arthaud Day et al. 2006; Collins et al., 2009; Leone and Liu, 2010), larger impairments in year $t-1$ are likely to reflect such circumstances. All other years are not statistically different from years surrounding CEO turnover. As an extra additional test, we eliminate all CEO turnover years. The results (not tabulated show) show that both the likelihood and size of asset impairments in the year following a CFO appointment are greater, consistent with Geiger and North (2006) and with results for internal CFO replacements in Table 6.¹⁶ Overall, results of these additional tests confirm that CFOs can exert independent and additional influence over financial reporting choices surrounding turnover, similar to that of CEOs.

5.4.3 Other impairment sensitivity tests

We undertake several other additional tests (not tabulated), to confirm our results. First, we partition the sample according to whether there is a change in CFO qualification (i.e., qualified accountant) arising from CFO turnover.¹⁷ The results for the sample of firms in which there is no change in qualification show lower likelihood and size of impairment in the year prior to turnover, consistent with the results reported in Table 5. This is unsurprising given that there is little variation in qualification within our sample (as shown in Panel C of Table 3). When testing instances in which the outgoing CFO is not qualified and is replaced with a qualified CFO ($n=130$), or vice versa ($n=83$), we find no significant results and the regressions are largely insignificant due to the small sample sizes. When we combine all changes in CFO qualifications ($n=213$) we also find no results. Therefore, qualification doesn't appear to impact on the impairment choice.

Second, we attempt to analyse the effect of CFO equity incentives on impairment decisions around CFO turnover by focusing on a sub-group of the sample where the incoming CFO's pay package includes an equity reward mechanism that is not in place for the outgoing CFO ($n=102$). The results show a significant decrease in the likelihood and size of impairments in year $t-1$, and a significant increase in impairment likelihood in year $t+1$. We also analyse the

¹⁶ Unreported tests indicate that the significant increase in impairments in year $t+1$ for CFO-only appointments is not justified by indicators of impairment.

¹⁷ "Qualified" defined as possession of a professional accounting certification (i.e., CA, CPA, CIMA, IPA or ACCA).

subsample in which the outgoing CFO's pay package includes an equity reward mechanism that is not in place for the incoming CFO (n=135). The results show a significant decrease in the likelihood and size of impairments in year $t-1$, and a significant increase in impairment in likelihood and size in year t .¹⁸ Overall, the results suggest that fewer (higher) asset impairments in year $t-1$ (t and $t+1$) occur regardless of whether the outgoing CFO has equity incentives. It is therefore reasonable that other concerns (e.g., career) in addition to equity reward mechanisms, serve as significant motivating factors in departing CFOs' financial reporting decisions.

Third, we investigate if a change in external auditor in the year of CFO turnover is associated with the likelihood or size of impairments. The findings indicate that in firms with a concurrent CFO and auditor change (n=293) there is an insignificant effect of CFO turnover on asset impairment likelihood and size. In contrast, the results for firms without an auditor change are consistent with those for the full sample reported in Table 5. These results suggest that a new auditor may mitigate earnings management in the year of a CFO appointment.¹⁹

5.4.4 CFO turnover and impairments relative to all firm-years

This study examines whether there is a change in the likelihood and size of non-current asset impairments surrounding CFO turnover, hence the sample is constructed from new CFO appointments. An investigation of the effect of CFO turnover on asset write-offs more broadly requires non-current asset impairment data for all ASX listed entities. This examination is problematic because non-current asset impairments are not electronically available and need to be hand collected from annual reports. Moreover, due to the frequency of CFO turnover, it is difficult to isolate firm-years, which are impervious to CFO turnover events occurring within neither the year prior to ($t-1$) nor immediately after ($t+1$) a CFO change.

Within the bounds of these limitations, we perform additional analyses by combining the test sample with 1,379 firm-year observations for Top 500 ASX listed entities for the period

¹⁸ For both subgroups, further testing suggests these significant results on impairment likelihood and size are not related to indicators of impairment.

¹⁹ Only 30.6 percent of firms with an auditor switch from a non-Big 4 to a Big 4 audit firm and, therefore, this result cannot be purely attributed to a "Big 4 audit quality" effect.

2007 to 2015.²⁰ Model (1) is then estimated for this expanded sample with the results presented in Table 9.

[Insert Table 9 about here]

The evidence reported in Table 9 indicates that a new CFO appointment is significantly positively associated with the likelihood (odds ratio = 1.421, $p < 0.05$) and size ($\beta = 0.083$, $p < 0.01$) of non-current asset impairments. Asset impairments are also more likely and increase around CEO turnover (*CEOTURNOVER*) The results for *BM* are significant and in line with predictions, however the two other external indicators of impairment (*BHR* and *YRS*) are insignificant, apart from *BHR* which is positive and significantly related to the likelihood of impairment. As expected, *BIG4* and *LMVE* are also positive and significantly related with asset impairment likelihood and size. Due to the concern that the merged sample may include firm-years before (t-1) or after (t+1) a CFO change, only CFO (i.e., year t) is included as the key test variable in the model. In additional analyses (not tabulated) we include *CFOTURN_PY*, *CFOTURN* and *CFOTURN_FY* in the model simultaneously and document positive and significant relations between *CFOTURN* and *CFOTURN_FY* and *IMPAIR* and *IMPAIRSIZE*.

5.4.5 CFO turnover and discretionary accruals

Given the absence of a single comprehensive proxy for accounting quality, the hypotheses are also tested by examining the effect of CFO turnover on signed performance-adjusted discretionary accruals (*PDA*). Financial data for all ASX listed companies between 2007 and 2016 are retrieved from the *Morningstar DatAnalysis Premium* database. This data is used to estimate the cross-sectional, performance-matched, modified Jones (1991) model (DeFond and Jiambalvo, 1994; Kothari et al., 2005).²¹ Expected accruals are estimated cross-sectionally for this sample of firm-years based on year groupings using the following model:

$$TA_{i,t}/A_{i,t-1} = \alpha_1(1/A_{i,t-1}) + \alpha_2(\Delta REV_{i,t} - \Delta REC_{i,t}/A_{i,t-1}) + \alpha_3(PPE_{i,t}/A_{i,t-1}) + \varepsilon \quad (2)$$

²⁰ The additional observations are limited to Top 500 ASX firms to minimize data collection requirements. Within the extra firm-years, firm-years in which a CEO, CFO or concurrent CEO and CFO appointment occur are excluded. Other deletions are consistent with the sample selection criteria in this study and pertain to delisted firms; firms listed in the current year; firms with a financial year-end other than 30 June; financial and utilities firms; and firms which apply fair value accounting.

²¹ After deleting observations with missing variables and firms in the consumer staples industry applying *AASB 141 Agriculture* and real-estate investment (applying *AASB 140 Investment Property*) sectors the final sample is 9,423 firm-years.

where TA (total accruals) is defined as net profit (loss) after tax before special items less cash flow from operations for firm i in period t .

The estimated coefficients from the yearly regressions of model (2) are used to calculate the value of expected accruals for sample firms. Discretionary accruals (DA) are defined as the difference between total accruals and expected accruals. The firm's estimated discretionary accruals ($DA_{i,t}$) are adjusted for performance matching following the approach used in prior research (Barua et al., 2010; Francis et al., 2005; Geiger and North, 2011). Performance decile groups by industry are formed based on the current year's ROA . Performance-adjusted discretionary accruals (PDA) are estimated as the difference between $DA_{i,t}$ and the median unadjusted DA for the industry ROA decile to which firm i belongs (where the median is calculated excluding firm i). Year groupings (as opposed to two-digit GICs industry classification and year groupings) are used to improve the significance and explanatory power of the models. Kothari et al. (2005) acknowledge the imprecision of the regression model based discretionary accrual estimates for observations with a limited number in any two-digit SIC code in any given year.²²

The following equation is estimated to examine the association between CFO turnover and PDA . The model includes controls for factors found to relate to discretionary accruals in prior research:

$$PDA = \beta_0 + \beta_1 CFOTURN + \beta_2 LNMVE + \beta_3 BME + \beta_4 CFFO + \beta_5 CR + \beta_6 LEVERAGE + \beta_7 PROA + \beta_8 LOSS + \beta_9 GOV_NED + \beta_{10} BIG4 + \varepsilon \quad (3)$$

Table 10 presents the results from estimating regression model (3).

[Insert Table 10 about here]

The results provide evidence of greater income increasing discretionary accruals in year $t+1$ after CFO turnover. To ensure discretionary accruals are not substituting for impairments, we also run our tests after including PDA as a control variable in our original models. The results (not tabulated) show that PDA reduces the likelihood of impairment across all

²² The mean (median) PDA for the CFO turnover sample of 1,389 firm-years is -0.001 (-0.003) and it is evident that extreme estimates of unexpected accruals (i.e., > 10 percent of prior total assets) are present. Consistent with prior literature (e.g., Kothari et al., 2005; Wells, 2002), the results indicate considerable imprecision in estimating unexpected accruals using the modified Jones (1991) model.

subsamples, while the coefficients on *CFOTURN_PY*, *CFOTURN* and *CFOTURN_FY* are consistent with our main results across all subsamples.²³

6. Conclusions

Based on a sample of 463 Australian CFO changes between 1 July 2007 to 30 June 2015 (i.e., 1,389 firm-years), this study examines the effect of CFO turnover on the likelihood and size of non-current asset impairments. The evidence indicates that *incoming* CFOs often initiate large asset write-offs in the initial or subsequent period of their appointment. The effect is largest for external hires with prior listed experience, and those receiving equity-based compensation. Therefore, the financial reporting motives maintained by CFOs surrounding a turnover event are similar to those of CEOs.

For *outgoing* CFOs, the results indicate that they are less likely to recognise (and record smaller) asset write-offs in the year prior to their departure. These findings are consistent with outgoing CFOs acting in self-interest, suggesting that CFOs are motivated to manage earnings upward to maximise their final pay (e.g., Geiger and North, 2006) or enhance their reputation before joining another firm. Prior research indicates that the value of CFOs' human capital in a competitive managerial market is influenced by the performance of the CFO's firm (Brickley et al., 1999; Fama, 1980; Mian, 2001). Our results support this view, suggesting that CFOs' motives to pursue post-retirement and other executive progression career prospects may also serve as an incentive to manipulate income prior to their departure.

The results in this study indicate that CFOs' incentives to manipulate earnings through asset write-offs exist irrespective of concurrent CEO turnover. This evidence demonstrates that CFOs independently influence the timing and magnitude of asset write-offs to bias earnings surrounding a turnover event. As the impairment of non-current assets continues to be a focus area of ASIC's inquiries, these results suggest that heightened scrutiny should be directed at firms appointing a new CFO. The findings are also relevant to the impairment literature, indicating that CFO turnover should be considered a key determinant of non-current asset impairments in addition to CEO turnover.

²³ The correlation between impairment and DA and PDA is low (-0.1457 and -0.1536 respectively).

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Data availability: The data that support the findings of this study are available from Connect4, SIRCA and Morningstar DatAnalysis Premium databases. Restrictions apply to the availability of these data, which were used under license for this study. Data regarding classification of CFO turnovers as routine and nonroutine are available from the authors upon request.

Table 1: Sample Construction

	CFO Changes	Firm-Years
New CFO appointments by ASX listed entities from 1 July 2007 to 30 June 2015:	2,602	7,806
<i>Less Exclusions:</i>	<i>2,139</i>	<i>6,417</i>
CFO turnover events in delisted entities, entities in administration, and firms which are newly listed in the year of or prior to the CFO appointment (i.e. financial years t and $t-1$ respectively):	701	2,103
CFO changes in firms which have a financial year end other than 30 June:	337	1,011
CFO appointments by entities in the banking, insurance, utilities, diversified financials or real estate industries; firms whose principal activities comprise of the management/harvest of biological assets; or firms identified as carrying items of PPE at fair value:	243	729
Cases where the appointment date is merely a position change for the incumbent CFO who absorbs the Company Secretary function:	241	723
Multiple CFO appointments within the same financial year (t), or with consecutive CFO appointments in subsequent financial years:	449	1,347
Turnover events where the incoming CFO left within the year of appointment (t) or the following financial year ($t+1$):	139	417
Firms with CFO changes that are only one financial year apart giving rise to duplicate firm-years in the sample:	29	87
Total usable observations:	463	1,389

Table 2: Reason for CFO departure

	CFO Changes	Firm-Years
Unique CFO appointments during the sample period 1 July 2007 to 30 June 2015:	463	1,389
<i>Reason for departure identified through expanded information search (i.e., Financial Statements, ASX Announcements, News Articles):</i>	33	99
Former CFO made redundant or CFO change follows the announcement of a formal restructure of the board or senior executive leadership team:	17	51
Resignation of former CFO after serving in an Acting CEO/MD position or CFO reported as being overlooked for the CEO/MD position:	2	6
CFO change as a result of the ill health or death of the outgoing CFO:	3	9
Reason for CFO departure consistently described as a retirement and outgoing CFO is above (or close to) the natural retirement age of 65:	5	15
Outgoing CFO resignation disclosed as being due to personal reasons, family commitments (e.g., maternity leave) or an extended career break:	6	18
Circumstances of change not disclosed in annual report and could not be identified by reviewing other information channels:	430	1,290

Table 3: Descriptive Statistics

<i>Variables</i>	<i>N</i>	<i>Mean</i>	<i>Median</i>	<i>Std. Dev</i>	<i>Min</i>	<i>Max</i>
Panel A: Asset Impairment						
<i>IMPAIR</i>	1,389	0.270	0.000	0.444	0.000	1.000
<i>IMPAIRSIZE</i>	1,389	0.015	0.000	0.041	0.000	0.162
Panel B: Control Variables						
<i>BM</i>	1,389	1.041	0.697	0.948	0.107	3.631
<i>YRS</i>	1,349	0.193	0.000	0.395	0.000	1.000
<i>BHR</i>	1,389	0.073	-0.106	0.749	-0.781	2.158
<i>CF</i>	1,389	0.022	-0.009	0.166	-0.259	0.529
<i>EARN</i>	1,389	0.071	-0.004	0.195	-0.092	0.714
<i>LN MVE</i>	1,389	10.916	10.573	2.068	7.905	15.323
<i>GOV_NED</i>	1,389	0.842	1.000	0.365	0.000	1.000
<i>BIG4</i>	1,389	0.437	0.000	0.496	0.000	1.000
<i>CEO_TURNOVER</i>	1,389	0.158	0.000	0.365	0.000	1.000
Panel C: CFO Turnover Characteristics						
<i>I_CFO_EXTERNAL</i>	1,386	0.784	1.000	0.412	0.000	1.000
<i>I_CFO_PRIOR_LISTED</i>	1,017	0.661	1.000	0.474	0.000	1.000
<i>I_CFO_QUALIFIED</i>	1,356	0.869	1.000	0.337	0.000	1.000
<i>I_CFO_EQUITY</i>	1,284	0.750	1.000	0.433	0.000	1.000
<i>O_CFO_CONTINUE</i>	987	0.219	0.000	0.414	0.000	1.000
<i>O_CFO_NEW_LISTED</i>	627	0.694	1.000	0.461	0.000	1.000
<i>O_CFO_QUALIFIED</i>	942	0.815	1.000	0.388	0.000	1.000
<i>O_CFO_EQUITY</i>	930	0.794	1.000	0.405	0.000	1.000

Definitions of the variables are presented in Appendix A. All continuous variables have been winsorised at the 5 and 95 percentiles.

Table 4: Mean Comparison of variables (*t*-test) by year around CFO turnover

<i>Variables</i>	<i>Mean</i>			<i>t-test</i>		
	<i>(t-1)</i>	<i>(t)</i>	<i>(t+1)</i>	<i>t vs t-1</i>	<i>t+1 vs t-1</i>	<i>t vs t+1</i>
Panel A: Asset Impairment						
<i>IMPAIR</i>	0.216	0.300	0.294	0.084***	0.078***	0.006
<i>IMPAIRSIZE</i>	0.011	0.019	0.016	0.008***	0.005**	0.003
Panel B: Control Variables						
<i>BM</i>	0.918	1.064	1.141	0.146**	0.223***	-0.077
<i>YRS</i>	0.172	0.188	0.218	0.016	0.046*	-0.030
<i>BHR</i>	0.147	0.025	0.046	-0.122**	-0.101**	-0.021
<i>EARN</i>	0.073	0.072	0.068	-0.001	-0.005	0.004
<i>CF</i>	0.018	0.022	0.025	0.003	0.006	-0.003
<i>BIG4</i>	0.413	0.438	0.460	0.026	0.048	-0.022
<i>GOV_NED</i>	0.819	0.849	0.860	0.030	0.041*	-0.011
<i>LNMVE</i>	10.959	10.920	10.870	-0.038	-0.089	0.051

Definitions of the variables are presented in Appendix A. All continuous variables have been winsorised at the 5 and 95 percentiles.

Table 5: Non-Current Asset Impairments

Variables	Predicted	<i>IMPAIR</i>			<i>IMPAIRSIZE</i>		
	Sign	(1)	(2)	(3)	(4)	(5)	(6)
<i>CFOTURN_PY</i>	?	0.717** (-2.009)			-0.062** (-2.054)		
<i>CFOTURN</i>	?		1.167 (1.029)			0.063** (2.246)	
<i>CFOTURN_FY</i>	?			1.159 (0.928)			-0.007 (-0.258)
<i>CEOTURNOVER</i>	+	1.990*** (3.262)	1.968*** (3.203)	2.088*** (3.409)	0.105*** (2.846)	0.094** (2.525)	0.109*** (2.924)
<i>BM</i>	+	1.750*** (4.432)	1.766*** (4.477)	1.755*** (4.459)	0.127*** (6.322)	0.129*** (6.468)	0.129*** (6.413)
<i>YRS</i>	+	1.457 (1.524)	1.469 (1.562)	1.461 (1.541)	-0.036 (-0.848)	-0.036 (-0.840)	-0.034 (-0.804)
<i>BHR</i>	-	0.679*** (-2.605)	0.678*** (-2.630)	0.675*** (-2.635)	-0.003 (-0.135)	-0.003 (-0.134)	-0.004 (-0.178)
<i>EARN</i>	-	0.309 (-1.482)	0.305 (-1.512)	0.306 (-1.502)	-0.168 (-1.499)	-0.168 (-1.513)	-0.170 (-1.523)
<i>CF</i>	-	2.258 (1.317)	2.275 (1.332)	2.253 (1.327)	0.009 (0.087)	0.010 (0.098)	0.011 (0.115)
<i>GOV_NED</i>	+	0.841 (-0.670)	0.859 (-0.592)	0.862 (-0.576)	-0.005 (-0.110)	-0.003 (-0.075)	-0.001 (-0.012)
<i>BIG4</i>	+	1.634** (2.234)	1.658** (2.305)	1.633** (2.241)	0.099** (2.380)	0.102** (2.470)	0.102** (2.463)
<i>LMNVE</i>	+	1.336*** (3.765)	1.332*** (3.739)	1.335*** (3.778)	0.029** (2.410)	0.029** (2.387)	0.029** (2.388)
<i>Constant</i>	?	0.013*** (-4.622)	0.011*** (-4.792)	0.011*** (-4.815)	-0.628*** (-4.406)	-0.669*** (-4.702)	-0.650*** (-4.559)
Industry FE		YES	YES	YES	YES	YES	YES
Observations		1,349	1,349	1,349	1,349	1,349	1,349
Chi-squared		81.79	80.37	79.20	93.98	95.07	90.69
Prob > chi2		0.000	0.000	0.000	0.000	0.000	0.000

Definitions of the variables are presented in Appendix A. All continuous variables have been winsorised at the 5 and 95 percentiles. For *IMPAIR* model (1) is estimated using a random-effects panel logistic regression and the odds ratio is reported with robust z-statistics in parentheses. The sign of the z-statistic indicates whether the predictor variable is associated with higher or lower odds of the outcome occurring. For *IMPAIRSIZE*, model (1) is estimated using a Tobit random-effects regression and the coefficient is reported with z-statistics in parentheses. *, ** and *** indicates significance at the 10%, 5% and 1% level respectively.

Table 6: Non-Current Asset Impairments and Characteristics of CFO turnover**Panel A: Outgoing CFO**

	Continues in Another Capacity		Does Not Continue		Goes to Listed Firm		Received Equity Pay	
	<i>IMPAIR</i>	<i>IMPAIRSIZE</i>	<i>IMPAIR</i>	<i>IMPAIRSIZE</i>	<i>IMPAIR</i>	<i>IMPAIRSIZE</i>	<i>IMPAIR</i>	<i>IMPAIRSIZE</i>
<i>CFOTURN_PY</i>	0.634 (-1.064)	-0.031 (-1.039)	0.698* (-1.708)	-0.016* (-1.947)	0.507** (-2.359)	-0.054** (-2.452)	0.458* (-1.844)	-0.068* (-1.857)
<i>CFOTURN</i>	1.584 (1.213)	0.034 (1.130)	1.178 (0.851)	0.01 -1.317	0.987 (-0.050)	0.011 (0.516)	1.761 (1.570)	0.041 (1.256)
<i>CFOTURN_FY</i>	0.987 (-0.032)	-0.002 (-0.073)	1.190 (0.839)	0.005 -0.582	1.899** (2.341)	0.039* (1.855)	1.133 (0.321)	0.017 (0.515)
<i>Controls</i>	YES	YES	YES	YES	YES	YES	YES	YES
<i>Industry_FE</i>	YES	YES	YES	YES	YES	YES	YES	YES

Panel B: Incoming CFO

	Internal Hire		External Hire		Prior Listed Experience		Received Equity Pay	
	<i>IMPAIR</i>	<i>IMPAIRSIZE</i>	<i>IMPAIR</i>	<i>IMPAIRSIZE</i>	<i>IMPAIR</i>	<i>IMPAIRSIZE</i>	<i>IMPAIR</i>	<i>IMPAIRSIZE</i>
<i>CFOTURN_PY</i>	0.582 (-1.581)	-0.018 (-1.489)	0.747 (-1.505)	-0.044* (-1.671)	0.771 (-0.677)	-0.075 (-1.239)	0.884 (-0.618)	-0.042 (-1.071)
<i>CFOTURN</i>	0.896 (-0.285)	-0.004 (-0.337)	1.215 (1.144)	0.047* (1.938)	1.509 (1.395)	0.113** (2.003)	1.024 (0.130)	0.061* (1.648)
<i>CFOTURN_FY</i>	1.869* (1.675)	0.020* (1.765)	1.070 (0.358)	-0.008 (-0.335)	0.849 (-0.466)	-0.046 (-0.800)	1.094 (0.464)	-0.023 (-0.600)
<i>Controls</i>	YES	YES	YES	YES	YES	YES	YES	YES
<i>Industry_FE</i>	YES	YES	YES	YES	YES	YES	YES	YES

Definitions of the variables are presented in Appendix A. All continuous variables have been winsorised at the 5 and 95 percentiles. For *IMPAIR* model (1) is estimated using a random-effects panel logistic regression and the odds ratio is reported with robust z-statistics in parentheses. The sign of the z-statistic indicates whether the predictor variable is associated with higher or lower odds of the outcome occurring. For *IMPAIRSIZE*, model (1) is estimated using a Tobit random-effects regression and the coefficient is reported with z-statistics in parentheses. *, ** and *** indicates significance at the 10%, 5% and 1% level respectively.

Table 7: Non-Current Asset Impairments Surrounding Routine Vs Non-Routine Turnover**Panel A: Routine Turnover**

Variables	Predicted	<i>IMPAIR</i>			<i>IMPAIRSIZE</i>		
	Sign	(1)	(2)	(3)	(4)	(5)	(6)
<i>CFOTURN_PY</i>	?	0.676 (-1.589)			-0.015* (-1.743)		
<i>CFOTURN</i>	?		0.996 (-0.018)			-0.000 (-0.002)	
<i>CFOTURN_FY</i>	?			1.447 (1.576)			0.014* (1.693)
<i>CEOTURNOVER</i>	+	1.427 (1.116)	1.487 (1.261)	1.595 (1.444)	0.007 (0.562)	0.008 (0.645)	0.011 (0.910)
<i>BM</i>	+	2.250*** (4.199)	2.275*** (4.232)	2.240*** (4.173)	0.034*** (5.158)	0.034*** (5.257)	0.034*** (5.146)
<i>YRS</i>	+	1.067 (0.173)	1.090 (0.234)	1.074 (0.191)	-0.006 (-0.484)	-0.006 (-0.436)	-0.006 (-0.473)
<i>BHR</i>	-	0.638* (-1.747)	0.634* (-1.770)	0.630* (-1.742)	-0.009 (-1.177)	-0.010 (-1.215)	-0.009 (-1.193)
<i>EARN</i>	-	0.423 (-0.811)	0.411 (-0.853)	0.419 (-0.815)	-0.040 (-1.259)	-0.042 (-1.312)	-0.041 (-1.288)
<i>CF</i>	-	4.197* (1.784)	4.149* (1.772)	4.200* (1.781)	0.020 (0.729)	0.020 (0.744)	0.021 (0.759)
<i>GOV_NED</i>	+	0.921 (-0.169)	0.934 (-0.142)	0.958 (-0.089)	0.010 (0.629)	0.010 (0.657)	0.011 (0.724)
<i>BIG4</i>	+	2.437** (2.562)	2.466*** (2.602)	2.415** (2.527)	0.028** (2.111)	0.029** (2.170)	0.029** (2.123)
<i>LN MVE</i>	+	1.288** (2.126)	1.289** (2.164)	1.289** (2.138)	0.006 (1.487)	0.006 (1.550)	0.006 (1.497)
<i>Constant</i>	?	0.008*** (-3.150)	0.006*** (-3.259)	0.006*** (-3.316)	-0.150*** (-3.311)	-0.160*** (-3.531)	-0.162*** (-3.588)
Industry FE		YES	YES	YES	YES	YES	YES
Observations		611	611	611	611	611	611
Chi-squared		53.41	50.57	53.86	62.59	59.80	62.50
Prob > chi2		0.000	0.000	0.000	0.000	0.000	0.000

Definitions of the variables are presented in Appendix A. All continuous variables have been winsorised at the 5 and 95 percentiles. For *IMPAIR* model (1) is estimated using a random-effects panel logistic regression and the odds ratio is reported with robust z-statistics in parentheses. The sign of the z-statistic indicates whether the predictor variable is associated with higher or lower odds of the outcome occurring. For *IMPAIRSIZE*, model (1) is estimated using a Tobit random-effects regression and the coefficient is reported with z-statistics in parentheses. *, ** and *** indicates significance at the 10%, 5% and 1% level respectively.

Table 7: Non-Current Asset Impairments Surrounding Routine Vs Non-Routine Turnover**Panel B: Non-Routine Turnover**

Variables	Predicted	IMPAIR			IMPAIRSIZE		
	Sign	(1)	(2)	(3)	(4)	(5)	(6)
<i>CFOTURN_PY</i>	?	0.721 (-1.328)			-0.046 (-1.303)		
<i>CFOTURN</i>	?		1.380 (1.498)			0.073** (2.281)	
<i>CFOTURN_FY</i>	?			0.963 (-0.164)			-0.035 (-1.056)
<i>CEOTURNOVER</i>	+	2.718*** (3.454)	2.593*** (3.261)	2.745*** (3.383)	0.150*** (3.602)	0.137*** (3.290)	0.148*** (3.557)
<i>BM</i>	+	1.569*** (2.718)	1.580*** (2.724)	1.580*** (2.738)	0.064*** (2.878)	0.066*** (2.975)	0.066*** (2.976)
<i>YRS</i>	+	1.638 (1.420)	1.649 (1.419)	1.640 (1.406)	0.032 (0.668)	0.031 (0.638)	0.032 (0.669)
<i>BHR</i>	-	0.694* (-1.905)	0.694* (-1.912)	0.690* (-1.945)	-0.039 (-1.576)	-0.037 (-1.538)	-0.039 (-1.588)
<i>EARN</i>	-	0.154 (-1.590)	0.152 (-1.596)	0.154 (-1.591)	-0.256* (-1.791)	-0.256* (-1.805)	-0.253* (-1.773)
<i>CF</i>	-	0.956 (-0.045)	1.015 (0.015)	1.006 (0.006)	-0.054 (-0.381)	-0.043 (-0.311)	-0.043 (-0.304)
<i>GOV_NED</i>	+	0.784 (-0.745)	0.804 (-0.663)	0.817 (-0.615)	-0.034 (-0.726)	-0.030 (-0.652)	-0.027 (-0.576)
<i>BIG4</i>	+	1.432 (1.192)	1.465 (1.261)	1.443 (1.221)	0.064 (1.410)	0.068 (1.511)	0.068 (1.502)
<i>LN MVE</i>	+	1.366*** (3.122)	1.358*** (3.040)	1.362*** (3.092)	0.033** (2.313)	0.031** (2.233)	0.031** (2.209)
<i>Constant</i>	?	0.018*** (-3.400)	0.015*** (-3.536)	0.017*** (-3.484)	-0.531*** (-3.211)	-0.558*** (-3.392)	-0.525*** (-3.172)
Industry FE		YES	YES	YES	YES	YES	YES
Observations		738	738	738	738	738	738
Chi-squared		59.81	61.16	57.14	51.90	55.14	52.03
Prob > chi2		0.000	0.000	0.000	0.000	0.000	0.000

Definitions of the variables are presented in Appendix A. All continuous variables have been winsorised at the 5 and 95 percentiles. For *IMPAIR* model (1) is estimated using a random-effects panel logistic regression and the odds ratio is reported with robust z-statistics in parentheses. The sign of the z-statistic indicates whether the predictor variable is associated with higher or lower odds of the outcome occurring. For *IMPAIRSIZE*, model (1) is estimated using a Tobit random-effects regression and the coefficient is reported with z-statistics in parentheses. *, ** and *** indicates significance at the 10%, 5% and 1% level respectively.

Table 8: Comparison of Mean Impairment Size by Executive Turnover

	(1)	(2)	(2) – (1)	(3)	(3) – (1)	(4)	(4) – (1)
Year relative to turnover	CEO only	CFO only	Two-tailed t-test	CFO & CEO	Two-tailed t-test	CFO, CEO & Chair	Two-tailed t-test
Year t-1	0.010	0.020	0.778	0.003	-1.687*	0.027	2.131**
Year t	0.017	0.015	-0.385	0.025	1.077	0.022	0.896
Year t+1	0.018	0.026	0.7547	0.015	-0.359	0.025	0.9950

This table compares the mean impairment size in the year prior, during and following CFO (concurrent CEO/CFO and concurrent CEO/CFO/Chair) turnover to instances in which there is only CEO turnover. Variable definitions are presented in Appendix A. *, ** and *** indicates significance at the 10%, 5% and 1% level respectively.

Table 9: Non-Current Asset Impairments around CFO turnover relative to an expanded sample

Variables	Predicted Sign	<i>IMPAIR</i> (1)	<i>IMPAIRSIZE</i> (2)
<i>CFOTURN</i>	?	1.421** (2.417)	0.083*** (3.637)
<i>CEOTURNOVER</i>	+	2.022*** (3.704)	0.082*** (2.799)
<i>BM</i>	+	2.628*** (7.733)	0.158*** (9.401)
<i>YRS</i>	+	1.362 (1.619)	-0.018 (-0.609)
<i>BHR</i>	-	0.818* (-1.924)	-0.001 (-0.085)
<i>EARN</i>	-	0.591 (-1.167)	-0.078 (-1.302)
<i>CF</i>	-	1.095 (0.305)	-0.026 (-0.553)
<i>GOV_NED</i>	+	0.791 (-1.193)	-0.018 (-0.567)
<i>BIG4</i>	+	1.494** (2.252)	0.091*** (2.953)
<i>LMVE</i>	+	1.339*** (4.352)	0.028*** (2.928)
<i>Constant</i>	?	0.005*** (-6.344)	-0.735*** (-6.343)
Industry FE		YES	YES
Observations		2,768	2,768
Chi-squared		139	172.3
Prob > chi2		0.000	0.000

Definitions of the variables are presented in Appendix A. All continuous variables have been winsorised at the 5 and 95 percentiles. For *IMPAIR* model (1) is estimated using a random-effects panel logistic regression and the odds ratio is reported with robust z-statistics in parentheses. The sign of the z-statistic indicates whether the predictor variable is associated with higher or lower odds of the outcome occurring. For *IMPAIRSIZE*, model (1) is estimated using a Tobit random-effects regression and the coefficient is reported with z-statistics in parentheses. *, ** and *** indicates significance at the 10%, 5% and 1% level respectively

Table 10: CFO Turnover and Signed Performance Adjusted Discretionary Accruals

Variables	Predicted Sign	PDA (1)	PDA (2)	PDA (3)
<i>CFOTURN_PY</i>	?	-0.011 (-1.511)		
<i>CFOTURN</i>	?		-0.005 (-0.800)	
<i>CFOTURN_FY</i>	?			0.016** (2.140)
<i>CEOTURNOVER</i>	?	-0.022* (-1.790)	-0.019 (-1.565)	-0.018 (-1.500)
<i>LMVE</i>	+	0.017 (1.624)	0.018* (1.717)	0.017* (1.655)
<i>BME</i>	+	0.039*** (3.743)	0.042*** (4.039)	0.039*** (3.733)
<i>CFFO</i>	+	-0.432*** (-8.240)	-0.432*** (-8.202)	-0.432*** (-8.279)
<i>CR</i>	+	0.002 (1.245)	0.002 (1.176)	0.002 (1.285)
<i>LEVERAGE</i>	+	-0.165** (-1.977)	-0.156* (-1.859)	-0.165* (-1.955)
<i>PROA</i>	-	-0.036* (-1.726)	-0.039* (-1.828)	-0.035* (-1.679)
<i>LOSS</i>	-	-0.160*** (-12.367)	-0.160*** (-12.335)	-0.159*** (-12.325)
<i>GOV_NED</i>	?	0.002 (0.129)	0.004 (0.264)	0.002 (0.138)
<i>BIG4</i>	?	-0.011 (-0.445)	-0.009 (-0.344)	-0.013 (-0.519)
<i>Constant</i>	?	-0.146 (-1.183)	-0.164 (-1.345)	-0.158 (-1.296)
Observations		1,388	1,388	1,388
Adjusted R-squared		0.236	0.235	0.238
F-test		20.42	20.07	20.48
Prob > F		0.000	0.000	0.000

Definitions of the variables are presented in Appendix A. The models are estimated using a fixed-effects panel regression. Numbers in parentheses are robust t-statistics. *, ** and *** indicates significance at the 10%, 5% and 1% level respectively.

Appendix A: Definition of Variables

<i>Variables</i>	<i>Definition</i>
Panel A: Key Test Variables	
<i>CFOTURN</i>	A binary variable coded as one in the financial year of a new CFO appointment (i.e. t), zero otherwise
<i>CFOTURN_PY</i>	A binary variable coded as one in the financial year prior to a new CFO appointment (i.e. $t-1$), zero otherwise
<i>CFOTURN_FY</i>	A binary variable coded as one in the financial year following a new CFO appointment (i.e. $t+1$), zero otherwise
Panel B: Incoming and Outgoing CFO Variables	
<i>I_CFO_EXTERNAL</i>	A binary variable coded as one if the incoming CFO was appointed externally, zero otherwise.
<i>I_CFO_PRIOR_LISTED</i>	A binary variable coded as one if the incoming CFO has prior experience at a listed firm, zero otherwise.
<i>I_CFO_QUALIFIED</i>	A binary variable coded as one if the incoming CFO holds a professional accounting certification (e.g., CA, CPA, CIMA, IPA or ACCA), zero otherwise.
<i>I_CFO_EQUITY</i>	A binary variable coded as one if the incoming CFO receives equity-based pay, zero otherwise.
<i>O_CFO_CONTINUE</i>	A binary variable coded as one if the outgoing CFO continues at the firm in another capacity (e.g., as an independent director), zero otherwise.
<i>O_CFO_NEW_LISTED</i>	A binary variable coded as one if the outgoing CFO moves to another listed firm, zero otherwise.
<i>O_CFO_QUALIFIED</i>	A binary variable coded as one if the outgoing CFO holds a professional accounting certification (e.g., CA, CPA, CIMA, IPA or ACCA), zero otherwise.
<i>O_CFO_EQUITY</i>	A binary variable coded as one if the outgoing CFO received equity-based pay, zero otherwise.
Panel C: Asset Impairment variables	
<i>IMPAIR</i>	A binary variable coded as one if a non-current asset impairment is recognised, zero otherwise.
<i>IMPAIRSIZE</i>	The total non-current asset impairment loss (excluding any impairment reversals) recognised by the firm in the current financial year (t) scaled by $t-1$ total assets.
Panel D: Control Variables	
<i>BM</i>	The ratio of the total book value of equity (<i>adjusted for the recognition of asset impairments</i>) to market capitalisation.
<i>YRS</i>	A binary variable coded as one if $BM > 1$ or $BM < 0$ for two reporting periods prior to the current financial year end (i.e. $t-1$ and $t-2$), zero otherwise.
<i>BHR</i>	The buy-hold return for the firm's stock over the financial year, calculated as: $(P_1 - P_0) + \text{Dividends} / P_0$
<i>EARN</i>	Net profit after tax per share, adjusted for impairment charges.
<i>CF</i>	Aggregate cash flow from operating and investing per share.
<i>BIG4</i>	A binary variable coded as one if the company is audited by a Big 4 accounting firm, zero otherwise.
<i>GOV_NED</i>	A binary variable coded as one if the majority of board members are independent non-executive directors, zero otherwise.
<i>LN MVE</i>	The natural logarithm of the market capitalisation.
<i>CEOTURNOVER</i>	A binary variable coded one if a CEO appointment took place, zero otherwise.

Panel E: Discretionary accruals test

<i>PDA</i>	Signed performance-adjusted discretionary accruals estimated from the modified cross-sectional modified Jones (1991) model.
<i>BME</i>	The ratio of the total book value of equity to market capitalisation.
<i>CFFO</i>	Cash flow from operations divided by total assets.
<i>CR</i>	The ratio of current assets to current liabilities.
<i>LEVERAGE</i>	The ratio of book value of debt to book value of assets.
<i>LOSS</i>	A binary variable coded as one if the company reported a net loss after tax, zero otherwise.
<i>PROA</i>	Return on assets (defined as net profit after tax divided by total assets) for the prior period.
